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Comparis	son of two	samples	: overview
	x, y indep.?	Same sd.?	R
Example?	Yes	Yes	<pre>t.test(x,y, var.equal=T)</pre>
Example?	Yes	No	t.test(x,y)
Example?	No		<pre>t.test(x,y, paired=T)</pre>

For comparison of two groups some form of *t*-test may be used.

What about comparison of three or more groups? One-way ANOVA!





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lide 2 - Statistics for Life Science (Week 4-2) - Comparison of groups

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Antibiotics and the decomposition of organic material

Data

- Five types antibiotics and a control
- 36 heifers allocated to 6 treatment groups. Feed added antibiotics
- Dung suspended in bags in the ground
- Amount of organic material measured after 8 weeks.
- For spiramycin: only four measurements available

Problem:

• Does the antibiotics affect the decomposition of organic material?



Statistical model

Recall that g(i) denotes the group for observation *i*. For example

 $g(1) = \cdots = g(6) = \text{control}, \quad g(31) = \cdots = g(34) = \text{Spiramycin}$ $g(1) = \cdots = g(6) = 1, \qquad g(31) = \cdots = g(34) = 6.$

Statistical model: y_1, \ldots, y_{34} are independent and

 $y_i \sim N(\alpha_{g(i)}, \sigma^2)$

Parameters: $\alpha_1, \ldots, \alpha_6$ and σ .

Equivalently:

$$y_i = \alpha_{g(i)} + e_i, \quad e_1, \dots, e_{34} \sim N(0, \sigma^2)$$
 independent

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Estimation and confidence intervals Statistical model:

$$y_i = \alpha_{g(i)} + e_i, \quad e_1, \dots, e_n \sim N(0, \sigma^2)$$
 independent

Parameters: $\alpha_1, \ldots, \alpha_k$ and σ . In particular, we are interested in differences, $\alpha_i - \alpha_l$!

Estimates and standard errors:

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$$\hat{\alpha}_{j} = \bar{y}_{j}; \quad \text{SE}(\hat{\alpha}_{j}) = s\sqrt{1/n_{j}} = s/\sqrt{n_{j}}$$
$$\hat{\alpha}_{j} - \hat{\alpha}_{l} = \bar{y}_{j} - \bar{y}_{l}; \quad \text{SE}(\hat{\alpha}_{j} - \hat{\alpha}_{l}) = s\sqrt{1/n_{j} + 1/n_{l}}$$
$$\hat{\alpha} = s$$

Confidence intervals from the usual recipe:

estimate $\pm t_{0.975,n-k} \cdot \text{SE}(\text{estimate})$

NB. The pooled *s* is used, also when comparing two groups!

ė 3.0 Type ni Si Organic material 2.6 2.7 2.8 2.9 Control 6 2.603 0.119 2.895 0.117 α -cyperm. 6 Enrofloxacin 0.162 6 2.710 Fenbendaz. 6 2.833 0.124 lvermectin 6 3.002 0.1092.5 Spiramycin 4 2.855 0.054 2.4 Enr Fen Ive Alp Con Pooled estimate of the standard deviation: $s = \sqrt{\frac{1}{28} \left(5 \cdot s_1^2 + \dots + 3 \cdot s_6^2 \right)} = \sqrt{\frac{1}{34 - 6} \sum_{i=1}^n (y_i - \bar{y}_{g(i)})^2} = 0.1217$ ide 5 — Statistics for Life Science (Week 4-2) — Comparison of groups

Group means and standard deviations



Fitting the model:

One-way ANOVA in R

One-way ANOVA in R

Output from summary(model1):

Coefficients:

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	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	2.60333	0.04970	52.379	< 2e-16	***
<pre>factor(type)Alfacyp</pre>	0.29167	0.07029	4.150	0.000281	***
<pre>factor(type)Enroflox</pre>	0.10667	0.07029	1.518	0.140338	
factor(type)Fenbenda	0.23000	0.07029	3.272	0.002834	**
factor(type)Ivermect	0.39833	0.07029	5.667	4.5e-06	***
factor(type)Spiramyc	0.25167	0.07858	3.202	0.003384	**

Residual standard error: 0.1217 on 28 degrees of freedom

Interpretations:

- Estimate and CI for α_{cont} , $\alpha_{Fenb} \alpha_{cont}$ and α_{Fenb} ? Estimat for σ ?
- Why are the SE's not the same?

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Hypothesis. Variati groups	ion within and between
Hypothesis, $H_0: lpha_1 = \cdots =$ Alternative, $H_A:$ at least t	$= \alpha_k.$ wo α 's are different.
Organic material Organic material Organic material Organic material Organic material Organic Organi Organic Organic Organic Or	 Variation within groups — points around the lines SS_e = ∑ⁿ_{i=1}(y_i - ȳ_{g(i)})² Variation between groups — Lines around the dashed line SS_{grp} = ∑^k_{j=1} n_j(ȳ_j - ȳ)² Test statistic F = MS_{grp} = SS_{grp/(k-1)}/SS_e/(n-k)

C C					
> model1 <- lm(or > summary(model1)	g~factor((type))			
R chooses a reference g estimates changes comp	roup — the pared to tha	e first in alph at group.	nabetic or	[.] der — and	
We prefer the control gr	roup as the	reference gr	oup:		
> type <- relevel > model1 <- lm(or > summary(model1)	(type, re g~factor(ef="Contro] (type))	L")		
					all stop
					6
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(8				
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			1110021		211020
One-way ANOVA	in R.				
If we prefer to see the g	roup means	5:			
	<i>,</i> ,				
> model2 <- lm(org~fa	ctor(type)	-1)			
> summary(model2)					
coefficients:	Fatimato S	td Frror t		$Pr(\lambda +)$	
factor(type)Control	2 60333	0 04970	52 38	<20-16 ***	
factor(type)Alfacyp	2.89500	0.04970	58.25	<2e-16 ***	
factor(type)Enroflox	2.71000	0.04970	54.53	<2e-16 ***	
factor(type)Fenbenda	2.83333	0.04970	57.01	<2e-16 ***	
factor(type)Ivermect	3.00167	0.04970	60.39	<2e-16 ***	
factor(type)Spiramyc	2.85500	0.06087	46.90	<2e-16 ***	
Residual standard err	ar: 0.1217	on 28 dem	rees of f	freedom	
	01. 0.1217	on zo degi	.665 01 1	reedom	
					_
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Slide 11 — Statistics for Life Science (Week 4-2) — Compariso	on of groups				-

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Comparison af alle groupsne

Do not use model2 for this — only model1

Test statistic

$$F = \frac{\mathrm{MS}_{\mathrm{grp}}}{\mathrm{MS}_{e}} = \frac{\mathrm{SS}_{\mathrm{grp}}/(k-1)}{\mathrm{SS}_{e}/(n-k)}$$

Large values of F are in disagreement with the hypothesis. Hence, the p-value is

$$p = P(F \ge F_{obs}) = P(F \ge 7.97) = 0.00009$$

There is overwhelming evidence that the hypothesis is not true.

How did we get the *p*-value?

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Pairwise comparisons

Suppose we want to compare the control group (1) with the Fenbendazole group (4): $\alpha_4 - \alpha_1$.

Estimate and its standard error:

$$\hat{\alpha}_4 - \hat{\alpha}_1 = 2.833;$$
 SE $(\hat{\alpha}_4 - \hat{\alpha}_1) = 0.07029$

- Confidence interval for $\alpha_4 \alpha_1$?
- Test for the hypothesis $H_0: \alpha_1 = \alpha_4$?
- Do all the groups differ significantly from the control group?



The *F*-distribution

If the hypothesis is true, then the *F*-test statistic is *F*-distributed with (k-1, n-k) degrees of freedom.



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LSD-value: least significant difference A confidence interval for the difference $\alpha_i - \alpha_l$ is

$$\hat{\alpha}_j - \hat{\alpha}_l \pm \mathsf{LSD}$$

where

$$LSD_{j,l} = t_{0.975,n-k} \cdot SE(\hat{\alpha}_j - \hat{\alpha}_l) = t_{0.975,n-k} \cdot s \cdot \sqrt{1/n_j + 1/n_l}.$$

A t-test for the hypothesis that the difference is zero uses the test statistic

$$T = \frac{|\hat{\alpha}_j - \hat{\alpha}_l|}{\operatorname{SE}(\hat{\alpha}_j - \hat{\alpha}_l)}$$

which is *t*-distributed with n-k degrees of freedom. LSD for control and fenbend.: $2.048 \cdot 0.1217 \cdot \sqrt{1/6 + 1/6} = 0.144$

If all group sizes are the same, then so are the LSD-values:

$$\text{LSD} = t_{0.975, n-k} \cdot \text{SE}(\hat{\alpha}_j - \hat{\alpha}_l) = t_{0.975, n-k} \cdot s \cdot \sqrt{2/n'}$$

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Conclusion

Different effects of the different types has been shown with high degree of certainty (p < 0.0001)

For all types except Enrofloxacin the amount of organic material is significantly higher than for the control group.

These statements should be supplemented by estimates and confidence intervals for α 's and/or for differences to the control group.



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Summary: one-way ANOVA

- Statistical model: normal distribution with same SD in the groups; independence
- Estimation: group means and pooled SD
- Confidence interval: estimat $\pm t_{0.975,n-k} \cdot \text{SE}(\text{estimate})$
- Hypothesis of equal group means tested by $F = MS_{grp}/MS_e$.
- Pairwise comparisons conducted "within" the model, using all the observations to estimate the SD.

With only two groups, *t*-tests suffice. Different versions:

- Paired or unpaired?
- If unpaired: same SD or not?





Any time we make a test a type I error may occur. The risk depends on the level of significance — often 5%. One test: risk of type I error: 5% By *m* independent tests:

 $1 - 0.95^{m}$

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FACULTY OF LIFE SCIENCES Lecture summary: main points One-way ANOVA • Assumptions for one-way ANOVA • Hypotheses for one-way ANOVA • Test statistic and the *F*-fordelingen